

International Journal of Energy Economics and Policy

ISSN: 2146-4553

available at http: www.econjournals.com

International Journal of Energy Economics and Policy, 2021, 11(3), 155-162.



The Effect of the COVID-19 Pandemic on Stock Prices with the Event Window Approach: A Case Study of State Gas Companies, in the Energy Sector

Suripto¹*, Supriyanto²

¹Department of Business Administration, Faculty of Social and Political Science, Lampung University, Indonesia, ²Department of Business Administration, Faculty of Social and Political Science, Lampung University, Indonesia. *Email: suriptob.1969@fisip.unila.ac.id

Received: 01 December 2020 Accepted: 03 February 2021 DOI: https://doi.org/10.32479/ijeep.10999

ABSTRACT

Stock price data at State Gas Company is defined as the time-series data comprising varying volatility and heteroscedasticity. One of the best models used to solve the problem of heteroscedasticity is the GARCH (generalized autoregressive conditional heteroscedasticity) model. Therefore, this study aims to build the most suitable model for predicting the 186 days before and 176 days after the Covid-19 pandemic, as well as to provide recommendations to reduce the impact of daily stock price movements. Data were obtained by examining the daily stock price data in Indonesian National Gas Companies from 2019 to 2020. The study also discusses the Event Window, with the best model identified as AR (1) -GARCH (1,1). The result showed that an error of less than 0.0015 is AR (1) - GARCH (1,1), provided the best model for price forecasting of Indonesian National Gas Companies.

Keywords: Stock Price, Heteroscedasticity, GARCH Model, Event Window

JEL Classifications: C5, O42, Q4, Q47

1. INTRODUCTION

Forecasting is an estimation or prediction of a future occurrence by evaluating previous circumstances' information and data. Based on this instance, financial analysts as information mediators play an extensive role by examining useful data related to earnings and stock forecasts (Jahangir, 2013; Chunhui et al., 2013). They are also regarded as intermediaries because they carry out a retrospective analysis of the company's personal and financial information to predict future occurrences. Estimates made by financial analysts and the associated management aids to evaluate and assess companies as well as improve the quality of their financial reporting, which is a forecast of the expected revenue in the subsequent year (Beaver et al., 1980).

Forecasting is classified into three types of methods based on time, namely short, medium, and long term (Montgomery et al., 2008).

Short-term is adopted for daily, weekly, and monthly forecasting. Specifically, it aids the administration to make certain decisions regarding human resources, inventory control, and cash flow management (Fildes and Goodwin, 2007; Fama et al., 2005; Liu et al., 2020). Several studies relating to forecasting has been carried out, such as market models (Neslihanoglu et al., 2017), a country's recession, which is the major activity carried out by numerous economic institutions (Fornaro, 2016; Morana, 2017), volatility using the GARCH model (1,1) (Chia et al., 2016; Tsung-Han and Yu-Pin, 2013). The public presumes that volatility is similar to market risks.

The least stock price in the market is increased by volatility. Therefore, in order to realize capital gains, investors need to purchase these stocks as a long-term investment (Planning, National and Indonesia, 2020). The highest volatility depicts maximum uncertainties or returns. This situation is commonly referred to as the "Risk and Return Tradeoff."

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The COVID-19 pandemic did not only affect the health sector, it also eroded the global economy, including Indonesia (Baig et al., 2020), (Chen et al., 2020), (Just and Echaust, 2020), (Ortmann et al., 2020), (Singh, 2020). It affected the exchange rate, as well as caused a decline in the Composite Stock Price Index (IHSG), which eventually went into freefall. Furthermore, everything was beyond predictions and difficult to control. Prior to the confirmation of the first phase of COVID-19 in the country, the IHSG was at the level of 6244 (24 January), which was reduced to 5942 (20 February) and 5,361 (2 March). On March 12, when the WHO declared COVID-19, a global pandemic, the IHSG fell to 4.2 percent or 4937 during the Thursday session, a level that had not occurred in almost four years. Conversely, on March 13, stock trading was halted for the first time since 2008 due to the pandemic. (Planning et al., N.d., 2020)

In addition, all human activities were restricted in order to curb the spread of the virus. Several countries adopted partial and simultaneous restriction policies, which had an impact on energy demand.

Countries with full lockdown policies experienced lesser energy demand than those with partial lockdown rules. In 2020, a 6% decline was predicted in the previous year. This is presumed as the worst condition in 70 years after the second world war. Indonesia is one of the nations with limited restriction policies, which also impacted energy demand (Ibrahim et al., 2018).

However, supposing the daily volatility of energy is high, there tends to be either an enormous increase or decrease in stock price, thereby leading to the provision of trade benefits, which is referred to as "High-Risk High-Returns" (Hull, 2015; Zali et al., 2018; Lyócsa et al., 2020, (Ayinde et al., 2019). Investors that usually adopt strategic trading plans prefer high volatility (risk taker). On the contrary, those that tend to invest long-term prefers low volatility because stock prices are bound to increase in the future (risk of harm) (Chan and Wai-Ming, 2000; He et al., 2020; Lin et al., 2019). Several economic and statistical studies are currently used to predict market conditions (Dzikevičius and Šaranda, 2011; Gontijo et al., 2020).

Numerous studies have been carried out to discuss the effect of energy on economic growth and price forecasting. Tehran and

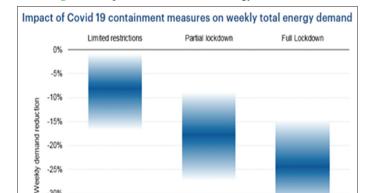


Figure 1: Impact of COVID-19 on energy demand world

Seyyedkolaee (2017), (Shinkevich et al., 2019) researched the relationship between oil price volatility and economic growth in Iran, an oil-exporting country. They also reviewed the impact of oil price volatility on domestic economic growth. Meanwhile, Vijayalakshmi et al. (2014) investigated the effect of price forecasts on the deregulated wholesale spot electricity market.

Weron and Misiorek (2006; 2008) studied the modeling of load forecasting and electricity prices. However, volatility in the stock market simply means the difference between an explosive increase or decrease in stock prices where there are moments when it goes up and down. Subsequently, when it is high, it implies that the stock price rises and falls significantly within one second. Volatility (price changes) in the capital market notably affects the return on investment. This circumstance also in accordance with risk and return trade-off theory known as "high-risk high-return." It is also considered as the basis for pricing assets and the acquisition of relevant information related to investment (Kongsilp and Mateus, 2017).

2. METHODOLOGY AND DATA

In this study, the data used was obtained from the stock price of State Gas Compan, the largest state-owned company in Indonesia. They are involved in the transmission and distribution of natural gas. Its business activities include planning, development, and management of downstream natural gas, processing, transportation, storage and trading, construction, production, as well as the supply, and distribution of artificial gas, etc (State Gas Company, 2020), (Ali et al., 2020), (Arafah et al., 2018), (Fadol, 2020), (Faizah and Husaeni, 2018), (Farhat et al., 2014), (Kapitonov and Voloshin, 2017).

The ability of the GARCH (p, q) model to fit properly is the main objective of this methodology. A brief introduction of this model and its equations, which are reported in full, before introducing the econometric considerations that need to be applied in this process are stated as follows.

2.1. Planning Data

The first stage of time series modeling is identification. It involves the calculation of ACF (autocorrelation function), PACF (partial autocorrelation function), and inverse autocorrelation from the time series data. Dickey and Fuller (1979) stated that supposing a distinction is required, it is relevant to carry out a stationary procedure.

2.2. Testing Stationary Data

The Augmented Dicky Fuller test (ADF) was used to evaluate stationary data, plot time-series graphs, and statistical analysis. However, some of the data tend to be non-stationary, such as price series, because they are not fixed. In addition, they are referred to as a unit-root non-stationary time series (Tsay, 2005). Unit-root is one of the features of certain stochastic processes that cause problems in time series modeling. The ADF test process is reported as follows (Brockwell and Davis, 2002; Tsay, 2005).

x1, x2 ..., xn are time series data and $\{xt\}$ follows the AR (p) model with mean μ . The model's mathematical expression is stated in equation (1).

-35%

$$Xt(\mu + \varphi 1Xt - 1) = \sum_{i=1}^{p-1} \varphi 1\Delta Xt - 1 + \varepsilon t \tag{1}$$

Where the difference in sequence xt, εt is white noise with 0 mean and variance $\sigma 2$ ($\varepsilon t \sim WN$ (0, $\sigma 2$)). The ADF analysis is a unit-root test that was realized by calculating the statistical value τ as follows:

Ho: $\phi 1 = 1$ (non-stationary data).

Ho: ϕ 1 < 1 (stationary data). Statistics.

Statistical test (ADF test):

$$\tau = \frac{\varphi 1}{Se\,\varphi 1} \tag{2}$$

Therefore, for the significance level ($\alpha = 0.05$), Ho is rejected supposing $\tau < -2.57$ or P < 0.05 (Brockwell and Davis, 2002).

2.3. Checking for White Noise

Subsequently, the use of a time series consisting of uncorrelated observations (data) has a constant variance, which is presumed to be white noise (Montgomery et al., 2008). On the contrary, when these time-series observations are normally distributed, it is referred to as the Gaussian white noise. Furthermore, when the time series is reported as white noise, the distribution of a large sample autocorrelation coefficient at lag k is similar to a normal distribution with 0 mean and a variance of 1/T, where T is the number of observations (Montgomery et al., 2008; Brockwell and Davis, 2002; Pankratz, 1991). The following expressions are reported in Equation (3).

$$r \sim N(0, T) \tag{3}$$

Based on Equation (3), it is possible to test the autocorrelation lag hypothesis k Ho: ρ k = 0 against Ha: ρ k \neq 0 by using the test statistics reported in Equation (4).

$$Z = \frac{rk}{\sqrt{1}/T} = rk\sqrt{T} \tag{4}$$

Ho is rejected when $|Z| > Z_{\alpha/2}^{\alpha}$ is on top of $\alpha/2$ percent of the standard or when P < 0.05. The test statistic realized from Equation (4) is used to evaluate the ACF and PACF (Wei, 2006). However, when the ACF is extremely slow decay, the time series is presumed to be non-stationary.

The aforementioned procedures are carried out, one at a time, specifically, the level of significance applies to autocorrelation and is considered individually. This study evaluates a set of autocorrelations together when the time series is reported as white noise. Therefore, to solve this problem, a statistical expression, adopted from the Box-Pierce statistic (Box-Pierce, 1970), was applied, as shown in Equation (5).

$$Q_{BP} = T \sum_{k=1}^{K} r^2 k \tag{5}$$

It is roughly distributed as chi-squared with degrees of freedom K, under the null hypothesis that the time series is white noise (Montgomery et al., 2008). Ho is rejected supposing Q BP>X

(a, K) 2 , it was concluded that the time series is not white noise. It is also possible to use the P-value in order to cause Ho to be rejected when P < 0.05.

Subsequently, supposing the data is not stationary, it becomes relevant to carry out the differentiation and transformation processes.

2.4. Testing the ARCH Effect

This step involves the estimation and examination of parameters, diagnoses, and test residuals, as well as selecting the best model based on certain criteria, such as determining the minimum value of AIC or SC. The residuals obtained from the best ARMA model were examined using the LM test to determine ARCH's effect. Although, when there is an ARCH effect, the data is modeled using the ARCH or GARCH method. The sequence of these models is discovered by plotting the square of the PACF residuals.

2.5. ARCH Model

The basic idea of the least square model assumes that the expected values for all squared errors are similar at some point, and this assumption is referred to as homoscedasticity (Engle, 2001). Meanwhile, the ARCH or GARCH model is based on the heteroscedasticity assumption that the variance is not constant. These models handle heteroscedasticity as a variant that needs to be modeled (Engle, 2001; Bollerslev, 1986). Engle (1982) introduced a time-variance model with an autoregressive conditional heteroscedasticity (ARCH) model using lagged disturbances. ARCH is an autoregression function that presumes that the variance is not constant over time and is also affected by previous data (Arch, 2006). The idea behind this model is to determine the relationship between the current and previous random variables.

2.6. Generalized ARCH (GARCH) Model

The GARCH (Generalized Autoregressive Conditional Heteroscedastic) model is a general form of ARCH. It was built to avoid an extremely high sequence. The GARCH model not only observes the relationship between several residuals, rather it also depends on some previous residuals (Eliyawati, 2014), and it was introduced by Bollerslev (1986), (Hsieh and Ritchken, 2005), (Virginia et al., 2018). The GARCH model with degrees p and q is defined as follows:

$$X_t \mid F_{(t-1)} \mid \sim N(0, \sigma_t^2)$$
 (6)

The GARCH model permits conditional variants based on previous lag, and this is reported in Equation (7).

$$\sigma_{t}^{2} = \omega + \sum_{i=1}^{q} \lambda_{i\varepsilon_{t-i}^{2}} + \sum_{j=1}^{p} \beta_{j\sigma_{t-j}^{2}}$$

$$\tag{7}$$

The present value of the conditional variant is parameterized based on the q and p lags of the squared residual and conditional variant. This is written as GARCH (p, q). Therefore, the conditional variance that varies from the GARCH model is heteroscedastic in accordance with the autoregression and MA (Wang, 2009). This model is reported in equation (8).

$$X_{t} = \delta + \sum_{i=1}^{p} \emptyset_{1} X_{t-i} - \sum_{i=1}^{q} \theta_{1} \varepsilon_{t-i} + \varepsilon_{t}$$
 (8)

 $\varepsilon_t \sim N(0, \sigma^2)$

$$\sigma_{t}^{2} = \omega + \sum_{i=1}^{q} \lambda_{i \varepsilon_{t-i}^{2}} + \sum_{j=1}^{p} \beta_{j \sigma_{t-j}^{2}}$$

x, is the equation of conditional mean (Bollerslev, 1986).

2.7. Model Selection Criteria

In selecting the ideal model, AIC criteria are used to discover the best predictions, and they are stated as follows:

$$AIC = -2\left(\frac{1}{T}\right) + 2\left(\frac{K}{T}\right),$$

Where

$$I = -\frac{Td}{2} (\ln 2\pi) - \frac{T}{2} \ln |\Omega|, |\Omega| = \det(\sum_{t} \varepsilon_{t} \varepsilon_{t}' / T)$$

l is the log-likelihood function, k is the number of parameters to be estimated, and T is the number of observations.

2.8. Checking the Event Window

Conceptually, the event window is the short-term deviation of a financial variable from its long-term level (Owens and Wu, 2011). The long and short-term levels depict the respective year and month sequentially. Therefore, the average year and month need to be calculated. In addition, the month's deviation from the mean of the year also needs to be discovered. Subsequently, the deviation is divided by the mean of the year and multiplied by 100 to determine the% deviation (Sahoo et al., 2012). Based on this concept, stock price behavior is compared to determine its average in a year.

3. RESULTS AND DISCUSSION

The data acquired from the stock price of State Gas Company before and after Covid-19 was utilized in this research. Before it was analyzed, a stationary data set was examined, and this was carried out in two ways, namely by (1) determining the data subjectivity plot and assessing whether or not the information is stationary (2) evaluating the stationary data using the ADF test.

The State Gas Company plot data is shown in Figure 1. The graph shows that the data is stationary, however three hundred and sixty-two of them portray an upward trend, which later moved downward to the final information. This behavior confirms that the data realized from the State Gas Company is constant at a certain number. Based on Table 1, the ADF unit-root test statistics for stationary data are reported in accordance with the test (P-value), which shows that the information acquired from the State Gas Company is 0.2097. It is, therefore ascertained that the data is stationary. Meanwhile, Table 2 shows that the test statistic for the intercept (Ho: Intercept = 0) is extremely significant with a P value> 0.0001. This means that its tapping is different from zero. In addition, the correlation analysis

of the data is shown in Figure 2. Based on these plots, there is a possibility of determining whether or not the State Gas Company data series is stationary. Therefore, the ACF indicates that the circuit is stationary because it decays extremely rapidly. Table 3 is used to determine the stationary data by checking WhiteNoise.

The White Noise behavior was used to check for data stationarity. This analysis is an approximate statistical test of the hypothesis, which indicates that there is no autocorrelation from the series to a specific break that is significantly different from zero. Although when this is true for all lags, then there is no information about the series. Autocorrelation was examined in six groups (Table 3) in which the hypothesis based on the white noise was strongly detected (P > 0.0001), which is to be expected because the State Gas Company data series (Figure 3) is stationary.

3.1. Identify the Different Series of Data State Gas Companies

Since the data series obtained is not stationary, the next step is to convert it to stationary using differentiation. Conversely, by

Table 1: Augmented Dickey-Fuller unit root test

| Type | Data | Lags | Tau | P-value |
|------|----------|------|--------|---------|
| Mean | PGAS Tbk | 2 | 0,9591 | 0,2097 |

Figure 2: State Gas Company data plot

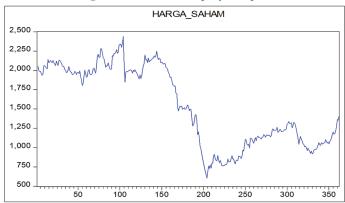


Figure 3: Correlation analysis of State Gas Company data

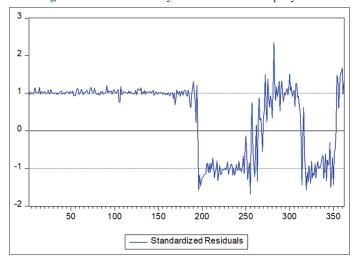


Table 2: Estimated parameters for tapping

| Variable | Data | DF | Estimate | Standard Error | t-value | P-value |
|-----------|----------|----|----------|----------------|---------|---------|
| Intercept | PGAS Tbk | 1 | 0.994 | 0.994 | 359.56 | 0.000 |

Table 3: Checking white noise on State Gas Company data

| Table 5: Checking white noise on State Gas Company da | | | | | | |
|---|----------|-------|--------|--------|-------|--|
| To lag | P-value | AC | Pac | Q-Stat | Prob. | |
| 1 | < 0,0001 | 0.994 | 0.994 | 359.56 | 0.000 | |
| 2 3 | < 0,0001 | 0.987 | -0.078 | 715.02 | 0.000 | |
| 3 | < 0,0001 | 0.981 | 0.110 | 1067.4 | 0.000 | |
| 4 | < 0,0001 | 0.975 | -0.048 | 1416.3 | 0.000 | |
| 5 | < 0,0001 | 0.968 | -0.028 | 1761.5 | 0.000 | |
| 6 | < 0,0001 | 0.961 | -0.075 | 2102.5 | 0.000 | |
| 7 | < 0,0001 | 0.953 | -0.013 | 2439.0 | 0.000 | |
| 8 | < 0,0001 | 0.945 | -0.064 | 2770.6 | 0.000 | |
| 9 | < 0,0001 | 0.937 | -0.032 | 3097.2 | 0.000 | |
| 10 | < 0,0001 | 0.928 | -0.029 | 3418.5 | 0.000 | |
| 11 | < 0,0001 | 0.919 | 0.005 | 3734.5 | 0.000 | |
| 12 | < 0,0001 | 0.910 | -0.002 | 4045.4 | 0.000 | |
| 13 | < 0,0001 | 0.901 | -0.015 | 4350.9 | 0.000 | |
| 14 | < 0,0001 | 0.892 | 0.056 | 4651.5 | 0.000 | |
| 15 | < 0,0001 | 0.884 | -0.026 | 4947.2 | 0.000 | |
| 16 | < 0,0001 | 0.875 | 0.041 | 5238.1 | 0.000 | |
| 17 | < 0,0001 | 0.866 | -0.047 | 5524.0 | 0.000 | |
| 18 | < 0,0001 | 0.857 | -0.061 | 5804.4 | 0.000 | |
| 19 | < 0,0001 | 0.846 | -0.086 | 6078.8 | 0.000 | |
| 20 | < 0,0001 | 0.836 | 0.036 | 6347.5 | 0.000 | |
| 21 | < 0,0001 | 0.826 | -0.024 | 6610.6 | 0.000 | |
| 22 | < 0,0001 | 0.817 | 0.077 | 6868.7 | 0.000 | |
| 23 | < 0,0001 | 0.808 | -0.038 | 7121.6 | 0.000 | |
| 24 | < 0,0001 | 0.797 | -0.035 | 7368.8 | 0.000 | |
| 25 | < 0,0001 | 0.788 | 0.050 | 7610.9 | 0.000 | |
| 26 | < 0,0001 | 0.778 | -0.037 | 7847.6 | 0.000 | |
| 27 | < 0,0001 | 0.767 | -0.058 | 8078.5 | 0.000 | |
| 28 | < 0,0001 | 0.756 | -0.051 | 8303.4 | 0.000 | |
| 29 | < 0,0001 | 0.745 | 0.012 | 8522.6 | 0.000 | |
| 30 | < 0,0001 | 0.734 | -0.005 | 8736.1 | 0.000 | |
| 31 | < 0,0001 | 0.723 | -0.046 | 8943.8 | 0.000 | |
| 32 | < 0,0001 | 0.712 | 0.009 | 9145.6 | 0.000 | |
| 33 | < 0,0001 | 0.701 | 0.059 | 9342.2 | 0.000 | |
| 34 | < 0,0001 | 0.691 | 0.027 | 9533.7 | 0.000 | |
| 35 | < 0,0001 | 0.681 | -0.006 | 9720.1 | 0.000 | |
| 36 | <0,0001 | 0.670 | -0.043 | 9900.9 | 0.000 | |

using the result from the differentiation as well as lag = 2 (d = 2), the State Gas Company data was obtained to be stationary. This is evident in residual data behavior after differentiation, which was approximately zero, as shown in Figure 3. Furthermore, this is also evident in the ACF plot's behavior, which was reported to decrease rapidly (Figure 3).

3.2. Testing the ARCH Effect

One of the key assumptions of ordinary least squares regression (OLS) is that the errors have similar variance (homoscedasticity). Although, when it is not constant across samples, the data is presumed to be heteroscedastic. This is because the OLS assumes constant variance, while the presence of heteroscedasticity makes its application inefficient for estimation. The models that take heteroscedasticity into account need to be applied to make the data more efficient. In regression analysis, a general linear model (GLM) is used to eradicate this issue. Conversely, during the time series analysis, several methods, such as the GARCH model, were applied. Therefore, before using this model, it is necessary to

Table 4: LM ARCH test data for State Gas Company

| Testing for ARCH interference based on OLS residue | | | | | | |
|--|-------------|--------------------|--------------|----------|--|--|
| Variable | Coefficient | Std. Error | t-Statistic | Prob. | | |
| С | 0.969870 | 0.060801 | 15.95163 | 0.0000 | | |
| $WGT_{}$ | 0.067770 | 0.052845 | 1.282444 | 0.2005 | | |
| $RESID^2(-1)$ | | | | | | |
| R-squared | 0.004573 | Mean depe | endent var | 1.040270 | | |
| Adjusted | 0.001792 | S.D. dependent var | | 0.496404 | | |
| R-squared | | | | | | |
| S.E. of regression | 0.495959 | Akaike info | o criterion | 1.440894 | | |
| Sum squared | 88.05921 | Schwarz criterion | | 1.462483 | | |
| resid | | | | | | |
| Log likelihood | -257.3608 | Hannan-Qı | iinn criter. | 1.449478 | | |
| F-statistic | 1.644662 | Durbin-W | atson stat | 1.994044 | | |
| Prob (F-statistic) | 0.200517 | | | | | |

check for the presence of heteroscedasticity, and the ARCH LM test is also be used.

Table 4 shows that the Q statistic is calculated based on the squared residual and is used to test for nonlinear effects (e.g., GARCH effect). The null hypothesis (Ho) is tested against Ha, as shown in Table 4:

Ho: OLS State Gas Company's residual data is white noise (or no ARCH effect was detected).

Against Ha: State Gas Company's OLS residual data is not white noise (or there is an ARCH effect).

Based on Table 5, it was discovered that AR (1) -GARCH (1,1) has a probability of 0.0070 and 0.0015. This is because the RMSE is extremely large, and this means that the model has better predictability. This is also supported by the forecasting and real value graph, which are extremely close (Figure 2). The Means Absolute Error (MAE) of 0.094 (Table 5) is also relatively small compared to the predicted stock price (H-1) (Table 6). The MAPE is 0.010 (Table 6), which is relatively small, indicating an ideal prediction accuracy.

In accordance with the Portmanteau Q test statistics and LM test, Ho was accepted because the P-value in Table 5 is P > 0.0001 (0.0015> 0.0001). It was therefore concluded that GARCH affects data acquired from the State Gas Company. This was also supported by the conditional variance behavior (Figure 3). Therefore, a model is needed to solve the issue of heteroscedastic variance. In this instance, the ARCH or GARCH model is used to explain the behavior of the data.

3.3. Windows Event Analysis

However, during the pandemic, from March to November, it was evident that the stock price was below its average in 2020. This has an AC value from the windows event test on the first day which was 0.917 till the 36th test when 0.101 was realized and kept

Table 5: GARCH State Gas Company estimated data statistics

| Testing the GARCH Estimate | | | | | | | |
|----------------------------|-------------|-------------|--------------|----------|--|--|--|
| | Coefficient | Std. Error | t-Statistic | Prob. | | | |
| С | 454.0191 | 246.7424 | 1.840053 | 0.0658 | | | |
| $RESID(-1)^2$ | 0.704713 | 0.261167 | 2.698325 | 0.0070 | | | |
| GARCH(-1) | 0.301634 | 0.094929 | 3.177472 | 0.0015 | | | |
| R-squared | -0.620968 | Mean depe | endent var | 1534.529 | | | |
| Adjusted R-squared | -0.620968 | S.D. deper | ndent var | 511.2199 | | | |
| S.E. of regression | 650.8711 | Akaike info | o criterion | 14.48315 | | | |
| Sum squared resid | 1.53E+08 | Schwarz | criterion | 14.52624 | | | |
| Log likelihood | -2610.209 | Hannan-Qı | uinn criter. | 14.50028 | | | |
| Durbin-Watson stat | 0.005715 | | | | | | |

Table 6: State Gas Company MAPE data statistics

| Variable | Coefficient | Std. Error | t-Statistic | Prob. |
|--|-------------|-------------|-------------|----------|
| HARGA SAHAM(-1) | 0.959112 | 0.010591 | 90.56202 | 0.0000 |
| D (HAR \overline{G} A SAHAM(-1)) | 0.135679 | 0.051696 | 2.624555 | 0.0091 |
| D (HARGA SAHAM(-2)) | -0.145159 | 0.051653 | -2.810270 | 0.0052 |
| C | 85.10233 | 22.21289 | 3.831214 | 0.0002 |
| INCPTBREAK | -41.47803 | 10.97470 | -3.779423 | 0.0002 |
| BREAKDUM | 37.57067 | 48.86662 | 0.768841 | 0.4425 |
| R-squared | 0.991423 | Mean depe | ndent var | 1530.531 |
| Adjusted R-squared | 0.991302 | S.D. deper | ndent var | 511.4746 |
| S.E. of regression | 47.70278 | Akaike info | criterion | 10.58447 |
| Sum squared resid | 800995.3 | Schwarz | criterion | 10.64951 |
| Log likelihood | -1888.621 | Hannan-Qu | inn criter. | 10.61034 |
| F-statistic | 8138.011 | Durbin-Wa | atson stat | 1.975206 |
| Prob (F-statistic) | 0.000000 | | | |

Table 7: The Average Abnormal Return Windows Event After Covid-19

| Checking Windows Events After Covid-19 | | | | | | |
|--|---------------------|----|-------|--------|--------|-------|
| Autocorrelation | Partial Correlation | | AC | PAC | Q-Stat | Prob* |
| . ****** | . ****** | 1 | 0.917 | 0.917 | 305.92 | 0.000 |
| ***** | . * | 2 | 0.873 | 0.203 | 583.99 | 0.000 |
| ***** | . ** | 3 | 0.876 | 0.343 | 865.06 | 0.000 |
| ***** | .[.] | 4 | 0.845 | -0.056 | 1126.9 | 0.000 |
| ***** | .[*] | 5 | 0.831 | 0.148 | 1380.8 | 0.000 |
| . ***** | * . | 6 | 0.801 | -0.147 | 1617.6 | 0.000 |
| . ***** | . . | 7 | 0.774 | 0.041 | 1839.1 | 0.000 |
| ***** | .j. j | 8 | 0.758 | -0.040 | 2052.4 | 0.000 |
| ***** | .j. j | 9 | 0.725 | -0.046 | 2248.0 | 0.000 |
| . ***** | .j. j | 10 | 0.700 | -0.014 | 2431.1 | 0.000 |
| . ***** | .j. j | 11 | 0.677 | -0.030 | 2602.6 | 0.000 |
| . ***** | * . | 12 | 0.637 | -0.096 | 2755.1 | 0.000 |
| **** | - i | 13 | 0.620 | 0.065 | 2900.0 | 0.000 |
| . **** | .j. j | 14 | 0.610 | 0.062 | 3040.4 | 0.000 |
| . **** | . * | 15 | 0.593 | 0.083 | 3173.5 | 0.000 |
| . **** | .j. j | 16 | 0.574 | -0.015 | 3298.7 | 0.000 |
| . **** | .j. j | 17 | 0.549 | -0.023 | 3413.6 | 0.000 |
| . **** | .j. j | 18 | 0.527 | -0.056 | 3519.9 | 0.000 |
| . **** | * . | 19 | 0.493 | -0.151 | 3612.8 | 0.000 |
| . *** | . . | 20 | 0.475 | 0.068 | 3699.7 | 0.000 |
| . *** | .j. j | 21 | 0.458 | -0.056 | 3780.6 | 0.000 |
| . *** | .j. j | 22 | 0.432 | 0.037 | 3852.8 | 0.000 |
| . *** | * . | 23 | 0.406 | -0.092 | 3916.8 | 0.000 |
| . *** | .i. i | 24 | 0.383 | 0.022 | 3973.8 | 0.000 |
| . *** | . * | 25 | 0.380 | 0.108 | 4030.0 | 0.000 |
| . ** | * . | 26 | 0.350 | -0.094 | 4077.9 | 0.000 |
| . ** | * . | 27 | 0.305 | -0.093 | 4114.5 | 0.000 |
| . ** | . . | 28 | 0.294 | 0.042 | 4148.6 | 0.000 |
| . ** | .j. j | 29 | 0.272 | -0.060 | 4177.7 | 0.000 |
| . ** | .j. j | 30 | 0.243 | -0.012 | 4201.1 | 0.000 |
| . ** | * . | 31 | 0.213 | -0.151 | 4219.1 | 0.000 |
| . * | . . | 32 | 0.188 | 0.061 | 4233.1 | 0.000 |
| . * | .j. j | 33 | 0.178 | 0.036 | 4245.8 | 0.000 |
| . * | .j. j | 34 | 0.156 | 0.045 | 4255.6 | 0.000 |
| . * | .j. j | 35 | 0.126 | -0.046 | 4262.0 | 0.000 |
| . * | * . | 36 | 0.101 | -0.082 | 4266.1 | 0.000 |

declining (Table 7). The percentage shows there is a possibility of a small event window due to the decline in stock price movements till December 2020.

4. CONCLUSION

In this study, the data from the State Gas Company of the Energy Sector was examined using the AR (p) -GARCH (p, q) time series analysis model. The results showed that the information is stationary. Furthermore, the differencing process was used with $\log = 2$ (d = 2) to convert the time series data to stationary. Conversely, by testing the effect of ARCH using the Q and LM tests, it was concluded that the GARCH model had an effect on the data realized from the State Gas Company. Based on this situation, AR (p) - GARCH (p, q) model was adopted.

The best model for the data acquired from State Gas Company is the AR (1) - GARCH (1,1) model. This is significant, and the R-squares are identified as 0.62 for the firm's model data. This prediction model's application is quite good based on the MAPE (the Mean Absolute Percentage Error) criterion for forecasting State Gas Company data that realized 0.094%. The model also needs to be used for forecasting in the next 176 days.

5. ACKNOWLEDGEMENT

The authors are grateful to the Ministry of Energy and Minerals and the Indonesia Stock Exchange for providing stock price data.

REFERENCES

- Ali, H.H., Abu Al-Rub, F.A., Shboul, B., Al Moumani, H. (2020), Evaluation of near-net-zero-energy building strategies: A case study on residential buildings in Jordan. International Journal of Energy Economics and Policy, 10(6), 325-336.
- Arafah, W., Nugroho, L., Takaya, R., Soekapdjo, S. (2018), Marketing strategy for renewable energy development in Indonesia context today. International Journal of Energy Economics and Policy, 8(5), 181-186.
- Arch, A. (2006), Analisis ARCH dan GARCH menggunakan. Eviews, 9, 1-17.
- Available from: https://www.coaction.id/dampak-covid-19-terhadap-permintaan-energi-dunia.
- Ayinde, A.R., Celik, B., Gylych, J. (2019), Effect of economic growth, industrialization, and urbanization on energy consumption in Nigeria: A vector error correction model analysis. International Journal of Energy Economics and Policy, 9(5), 409-418.
- Azimli, A. (2020), The impact of COVID-19 on the degree of dependence and structure of risk-return relationship: A quantile regression approach. Finance Research Letters, 36, 101648.
- Baig, A.S., Butt, H.A., Haroon, O., Rizvi, S.A.R. (2020), Deaths, panic, lockdowns and US equity markets: The case of COVID-19 pandemic. Finance Research Letters, 36, 101701.
- Beaver, W.H., Lambert, R., Morse, D. (1980), The information content of security prices. Journal of Accounting and Economics, 2, 3-28.
- Bollerslev, T. (1986), Generalized autoregressive conditional heteroskedasticity. Journal of Econometrics, 31, 307-327.
- Box, G.E.P., Pierce, D.A. (1970), Distribution of residual autocorrelations in autoregressive-integrated moving average time series models.

- Journal of American Statistics Association, 65, 1509-1526.
- Brockwell, P.J., Davis, R.A. (2002), Introduction to Time Series and Forecasting. New York: Springer-Verlag.
- Brooks, C. (2014), Introductory Econometrics for Finance. 3rd ed. New York: Cambridge University Press.
- Chan, K., Wai-Ming, F. (2000), Trade size, order imbalance, and the volatility-volume relation. Journal of Financial Economics, 57(2), 247-73.
- Chen, S., Yang, Y., Lin, J.H. (2020), Capped borrower credit risk and insurer hedging during the COVID-19 outbreak. Finance Research Letters, 36, 101744.
- Chia, C.L., Skindilias, K., Karathanasopoulos, A. (2016), Forecasting latent volatility through a Markov chain approximation filter. Journal of Forecasting, 35, 54-69.
- Choi, S.Y. (2020), Industry volatility and economic uncertainty due to the COVID-19 pandemic: Evidence from wavelet coherence analysis. Finance Research Letters, 37, 101783.
- Chunhui, L., Grace, O., Kwok-Kee, W., Lee, J. (2013), Ratio analysis comparability between Chinese and Japanese firms. Journal of Asia Business Studies, 7(2), 185-199.
- Ciner, C. (2020). Stock return predictability in the time of COVID-19. Finance Research Letters, 36, 101705.
- Dzikevicius, A., Šaranda, S. (2011), Smoothing techniques for market fluctuation signals. Business: Theory and Practice, 12(1), 63-74.
- Eliyawati, W. (2014). Penerapan model garch (generalized autoregressive conditional heteroscedasticity) Untuk menguji pasar modal efisien di Indonesia (studi pada harga penutupan (closing price) Indeks Saham LQ 45 Periode 2009-2011). Jurnal Administrasi Bisnis S1 Universitas Brawijaya, 7(2), 79049.
- Engle, R. (1982), Autoregressive conditional heteroscedasticity with estimates of the variance of United Kingdom inflation. Econometrica, 50, 987-1007.
- Engle, R. (2001), GARCH 101: The use of ARCH/GARCH models in applied econometrics. Journal of Economic Perspectives, 15(4), 157-168.
- Fadol, H.T.A. (2020), Study the possibility of address complex models in linear and non-linear causal relationships between oil price and GDP in KSA: Using the combination of toda-yamamoto, diks-panchenko and var approach. International Journal of Energy Economics and Policy, 10(6), 672-678.
- Faizah, S.I., Husaeni, U.A. (2018), Development of consumption and supplying energy in Indonesia's economy. International Journal of Energy Economics and Policy, 8(6), 313-321.
- Fama, E.F., Fisher, L., Jensen, M.C., Roll, R.W. (2005), The Adjustment of Stock Prices to New Information. International Economic Review, 10, 1-21.
- Farhat, R., Ghaddar, N.K., Ghali, K. (2014), Investing in PV systems utilizing savings from building envelop replacement by sustainable local material: A case study in Lebanese Inland region. International Journal of Energy Economics and Policy, 4(4), 554-567.
- Fildes, R., Goodwin, P. (2007), Against your better judgment? How organizations can improve their use of management judgment in forecasting. Interfaces, 37(6), 570-576.
- Fornaro, P. (2016), Forecasting U.S. recessions with a large set of predictors. Journal of Forecasting, 35, 477-492.
- Gontijo, J.B., Venturini, A.M., Yoshiura, C.A., Borges, C.D., Moura, J.M.S., Bohannan, B.J.M., Nüsslein, K., Rodrigues, J.L.M., Paula, F.S., Tsai, S.M. (2020), Seasonal dynamics of methane cycling microbial communities in Amazonian floodplain sediments. New York: bioRxiv.
- Gontijo, T.S., Rodrigues, A.C., De Muylder, C.F., La Falce, J.L., Pereira, T.H.M. (2020), Analysis of olive oil market volatility using the arch and garch techniques. International Journal of Energy

- Economics and Policy, 10(3), 423-428.
- Gunarto, T., Azhar, R., Tresiana, N., Supriyanto, S., Ahadiat, A. (2020), Accurate estimated model of volatility crude oil price. International Journal of Energy Economics and Policy, 10(5), 228-233.
- He, P., Sun, Y., Zhang, Y., Li, T. (2020), COVID-19's impact on stock prices across different sectors--An event study based on the Chinese stock market. Emerging Markets Finance and Trade, 56(10), 2198-2212.
- Hsieh, K.C., Ritchken, P. (2005), An empirical comparison of GARCH option pricing models. Review of Derivatives Research, 8(3), 129-150.
- Hull, J.C. (2015), Risk Management and Financial Institutions. Hoboken, NJ: John Wiley and Sons, Inc.
- Ibrahim, M.A., Myrna, R., Irawati, I., Kristiadi, J.B. (2018), Tax policy in indonesian energy sectors: An overview of tax amnesty implementation. International Journal of Energy Economics and Policy, 8(4), 234-236.
- Jahangir, A.K.A.M. (2013), Determinants and usefulness of analysts' cash flow forecasts: Evidence from Australia. International Journal of Accounting and Information Management, 21(1), 4-21.
- Just, M., Echaust, K. (2020), Stock market returns, volatility, correlation and liquidity during the COVID-19 crisis: Evidence from the Markov switching approach. Finance Research Letters, 37, 101775.
- Kapitonov, I.A., Voloshin, V.I. (2017), Strategic directions for increasing the share of renewable energy sources in the structure of energy consumption. International Journal of Energy Economics and Policy, 7(4), 90-98.
- Kongsilp, W., Mateus, C. (2017), Volatility risk and stock return predictability on global financial crises. China Finance Review International, 7(1), 33-66.
- Li, Y., Liang, C., Ma, F., Wang, J. (2020). The role of the IDEMV in predicting European stock market volatility during the COVID-19 pandemic. Finance Research Letters, 36, 101749.
- Lin, L., Hung, P.H., Chou, D.W., Lai, C.W. (2019), Financial performance and corporate social responsibility: Empirical evidence from Taiwan. Asia Pacific Management Review, 24(1), 61-71.
- Liu, H., Yi, X., Yin, L. (2020), The impact of operating flexibility on firms' performance during the COVID-19 outbreak: Evidence from China. Finance Research Letters, 36, 101808.
- Lyócsa, Š., Baumöhl, E., Výrost, T., Molnár, P. (2020), Fear of the coronavirus and the stock markets. Finance Research Letters, 36, 101735.
- Montgomery, D., Jennings, C., Kulachi, M. (2008), Introduction Time Series Analysis and Forecasting. Hoboken, NJ: John Wiley and Sons Inc
- Morana, C.J. (2017), The US dollar/euro exchange rate: Structural modeling and forecasting during the recent financial crises. Journal of Forecast, 36, 919-935.
- Neslihanoglu, S., Sogiakas, V., McColl, J.H., Lee, D. (2017), Nonlinearities in the CAPM: Evidence from developed and emerging markets. Journal of Forecast, 36, 867-897.
- Ortmann, R., Pelster, M., Wengerek, S.T. (2020), COVID-19 and investor behavior. Finance Research Letters, 37, 101717.
- Owens, E., Wu, J.S. (2011), Window Dressing of Financial Leverage. Available from: https://www.pdfs.semanticscholar.org/7669/

- e94d07af8aec27e068ae95d972a8b7b18098.pdf. [Last accessed on 2017 Nov 06].
- Pankratz, A. (1991), Forecasting with Dynamic Regression Models. Canada: Wiley Intersciences Publication.
- Perencanaan, K., Nasional, P., Indonesia, B.R. (2020), Dampak Covid-19 terhadap Pergerakan nilai tukar rupiah dan indeks harga saham gabungan (IHSG). Jurnal Perencanaan Pembangunan: The Indonesian Journal of Development Planning, 4(2), 151-165.
- PT Perusahaan Gas Negara Tbk. (2020), Available from: https://www.pgn.co.id.
- Sahoo, P.K., Chottray, R.K., Pattnaiak, S. (2012), Research issues on windows event log. International Journal of Computer Applications, 41(19), 40-48.
- Shinkevich, A.I., Kudryavtseva, S.S., Dyrdonova, A.N., Gallyamova, D.K., Farrakhova, A.A., and Vodolazhskaya, E.I. (2019), Assessment of the efficiency of energy-and resource-saving technologies in the model of open innovation. E3S Web of Conferences, 124(5), 289-296.
- Singh, A. (2020), COVID-19 and safer investment bets. Finance Research Letters, 36, 101729.
- Tehranchian, A.M., Seyyedkolaee, M.A. (2017), The impact of oil price volatility on the economic growth in Iran: An application of a threshold regression model. International Journal of Energy Economics and Policy, 7(4), 165-171.
- Tsay, R.S. (2005), Analysis of Financial Time Series. Hoboken, New Jersey: John Wiley and Sons, Inc.
- Tsung-Han, K., Yu-Pin, H. (2013), Forecasting volatility with many predictors. Journal of Forecasting, 32, 743-754.
- Vijayalakshmi, S., Ajaya, K.P., Badri, N.R. (2014), Forecasting electricity prices in deregulated wholesale spot electricity market: A review. International Journal of Energy Economics and Policy, 4(1), 32-42.
- Virginia, E., Ginting, J., Elfaki, F.A.M. (2018), Application of garch model to forecast data and volatility of share price of energy (Study on adaro energy Tbk,). International Journal of Energy Economics and Policy, 8(3), 131-140.
- Wang, P. (2009), Financial Econometrics. 2nd ed. New York: Routledge, Taylor and Francis Group.
- Wei, W.W. (2006), Time Series Analysis: Univariate and Multivariate Methods. 2nd ed. New York: Pearson.
- Weiss, A.A. (1984), ARMA models with ARCH errors. Journal of Time Series Analysis, 5, 129-143.
- Weron, R. (2006), Modeling and Forecasting Electricity Loads and Prices: A Statistical Approach. Chichester: Wiley Finance Publication.
- Weron, R., Misiorek, A. (2006), Short-term electricity price forecasting with time series models. A review and evaluation. In: Complex Electricity Markets: The European Power Supply Industry. Lodz: IEPL and SEP. p231-254.
- Weron, R., Misiorek, A. (2008), Forecasting spot electricity prices: A comparison of parametric and semiparametric time series models. International Journal of Forecasting, 24, 744-763.
- Zali, M., Yudi Heryadi, A., Nurlaila, S., Fanani, Z. (2018), Madura cattle agribusiness performance and feasibility in Galis region, Madura. International Journal of Advanced Multidisciplinary Research, 5(6), 45-55.